# Stochastic Models (Lecture #3)

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We would like to develop a concept of independence.

Intuitively, A and B are independent (notation:  $A \perp B$ ) if

$$\mathbb{P}[\boldsymbol{A}|\boldsymbol{B}] = \mathbb{P}[\boldsymbol{A}],$$

or, equivalently, if

$$\mathbb{P}[A]\mathbb{P}[B] = \mathbb{P}[A|B]\mathbb{P}[B] = \frac{\mathbb{P}[A \cap B]}{\mathbb{P}[B]}\mathbb{P}[B] = \mathbb{P}[A \cap B].$$

As a definition, we will say that

A and B are independent  $(A \perp B) \Leftrightarrow \mathbb{P}[A \cap B] = \mathbb{P}[A]\mathbb{P}[B]$ .

(since this also covers the cases where  $\mathbb{P}[B] = 0$ ).

Examples:

$$\mathcal{E}$$
: dice,  $\Omega = \{1, 2, \dots, 6\}, \ \mathcal{A} = \mathcal{P}(\Omega), \ \mathbb{P}[A] = \frac{\#A}{\#\Omega}.$ 

Are *A* = {6} and *B* = {4, 5, 6} independent? No, since

$$\frac{1}{6} \times \frac{1}{2} = \mathbb{P}[A] \times \mathbb{P}[B] \neq \mathbb{P}[A \cap B] = \frac{1}{6}.$$

► Are A ="obtain an even result" and B = {4, 5, 6} independent?

No, since

$$\frac{1}{2} \times \frac{1}{2} = \mathbb{P}[A] \times \mathbb{P}[B] \neq \mathbb{P}[A \cap B] = \frac{2}{3}.$$

Another example:

 $\mathcal{E}$ : 2 dices,  $\Omega = \{(1, 1), (1, 2), \dots, (6, 6)\}, \ \mathcal{A} = \mathcal{P}(\Omega), \ \mathbb{P}[A] = \frac{\#A}{\#\Omega}.$ 

- Are A ="obtain an even result for dice 1" and B ="obtain an even result for dice 2" independent? Yes, since

   \[P[A]P[B] = \frac{18}{36} × \frac{18}{36} = \frac{9}{36} = P[A ∩ B].
   \]
- Are A ="obtain an even result for dice 1" and C ="obtain an odd sum" independent? Yes, since

$$\mathbb{P}[A]\mathbb{P}[C] = \frac{18}{36} \times \frac{18}{36} = \frac{9}{36} = \mathbb{P}[A \cap C].$$

Are B ="obtain an even result for dice 2" and C ="obtain an odd sum" independent? Yes, since

$$\mathbb{P}[B]\mathbb{P}[C] = \frac{18}{36} \times \frac{18}{36} = \frac{9}{36} = \mathbb{P}[B \cap C].$$

But

In this example, A, B, C are pairwise independent.

# $\mathbb{P}[A]\mathbb{P}[B]\mathbb{P}[C] = rac{1}{2} imesrac{1}{2} imesrac{1}{2} eq 0 = \mathbb{P}[A\cap B\cap C].$

We will say that A, B, C are not mutually independent...

Definition: The events  $A_1, A_2, ..., A_n$  are (mutually) independent if and only if for all k, for all  $1 \le i_1 < i_2 < ... < i_k \le n$ , we have

$$\mathbb{P}[A_{i_1} \cap A_{i_2} \cap \ldots \cap A_{i_k}] = \mathbb{P}[A_{i_1}] \times \mathbb{P}[A_{i_2}] \times \ldots \times \mathbb{P}[A_{i_k}].$$

It is clear that the above example violates this definition for k = 3. Good, since, intuitively, we didn't want to claim that those 3 events are independent.

Let  $X_1, X_2, \ldots$  be r.v.'s on  $(\Omega, \mathcal{A}, P)$ .

$$\begin{array}{l} X_1, \dots, X_n \text{ are } \mathbb{L} \Leftrightarrow \\ \blacktriangleright \ F^X(x_1, \dots, x_n) = \prod_{i=1}^n F_{X_i}(x_i), \\ \blacktriangleright \ f^X(x_1, \dots, x_n) = f^{X_1}(x_1) \dots f^{X_n}(x_n) \text{ (provided existence...)} \end{array}$$

If 
$$X_1, \ldots, X_n$$
 are  $\bot$ ,  $\mathbb{E}[X_1 \ldots X_n] = \mathbb{E}[X_1] \ldots \mathbb{E}[X_n]$ .

**Corollary**:  $X \perp Y \Rightarrow Cov[X, Y] = 0$ .

Proof:  $X \perp Y \Rightarrow (X - \mathbb{E}[X]) \perp (Y - \mathbb{E}[Y])$ , so that  $Cov[X, Y] = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[X - \mathbb{E}[X]]\mathbb{E}[Y - \mathbb{E}[Y]] = 0$ .

Therefore, Cov[X, Y] can be considered as a measure of dependence between X and Y.

Remarks:

•  $\operatorname{Cov}[X, Y] = 0$  does not imply that  $X \perp Y$ . Example: *X*, with pdf *f*(.), symmetric about 0, and  $Y = X^2$ .

► 
$$\operatorname{Var}[X + Y] = \operatorname{Var}[X] + \operatorname{Var}[Y] + 2\operatorname{Cov}[X, Y]$$
, so that  $\operatorname{Var}[X + Y] = \operatorname{Var}[X] + \operatorname{Var}[Y]$  if  $X \perp Y$ .

Independence can be extended to  $\sigma$ -algebras:

Let  $A_1, A_2, \ldots \subset A$  be  $\sigma$ -algebras.

- $\blacktriangleright \ \mathcal{A}_1 \perp \mathcal{A}_2 \stackrel{\text{def}}{\Leftrightarrow} \forall A_1 \in \mathcal{A}_1, \forall A_2 \in \mathcal{A}_2, A_1 \perp A_2.$
- ▶  $\mathcal{A}_1, \ldots, \mathcal{A}_n$  are  $\mathbb{L} \stackrel{\text{def}}{\Leftrightarrow} \forall A_1 \in \mathcal{A}_1, \ldots, \forall A_n \in \mathcal{A}_n, A_1, \ldots, A_n$  are  $\mathbb{L}$ .
- ▶  $\mathcal{A}_1, \mathcal{A}_2, \dots$  are  $\mathbb{I} \stackrel{\text{def}}{\Leftrightarrow}$  for every  $n \ge 2, \mathcal{A}_1, \mathcal{A}_2, \dots, \mathcal{A}_n$  are  $\mathbb{I}$ .

Denote by

$$\sigma(X) \stackrel{\mathrm{def}}{=} X^{-1}(\mathcal{B})$$

the  $\sigma$ -algebra generated by the r.v. X. (See Lecture # 1.)

- Then  $X_1 \perp X_2 \Leftrightarrow \sigma(X_1) \perp \sigma(X_2)$ .
- This also allows for defining independence between r.v. and σ-algebras:

In the previous lecture, we studied how the concept of expectation can be used to compute the "most typical value" or "best guess" for some r.v. *X*.

Conditional expectation allows for exploiting some information (e.g., the information that some event occurred) in order to improve this (unconditional) best guess.

This information can take various forms:

- (the occurrence of) an event.
- (the value of ) a r.v.
- (the occurrence of some event in) a  $\sigma$ -algebra.

Let X be an integrable r.v. on  $(\Omega, \mathcal{A}, \mathbb{P})$ . Let  $A \in \mathcal{A}$ , with  $\mathbb{P}[A] \neq 0$ .

Then we let

$$\mathbb{E}[X|A] \stackrel{\text{def}}{=} \frac{1}{\mathbb{P}[A]} \int_A X(\omega) \, d\mathbb{P}(\omega) \left[ = \int_\Omega X(\omega) d\mathbb{P}(\omega|A) \right].$$

Remarks:

$$\blacktriangleright \mathbb{E}[X|A] = \frac{1}{\mathbb{P}[A]} \int_{\Omega} X(\omega) \mathbb{I}_{A}(\omega) d\mathbb{P}(\omega) = \frac{\mathbb{E}[X\mathbb{I}_{A}]}{\mathbb{P}[A]}.$$

► E[X|Ω] = E[X]. Intuitively, the information that Ω has occurred is void, so that the "initial best guess" E[X] cannot be improved.

• 
$$\mathbb{E}[\mathbb{I}_{A_1}|A_2] = \mathbb{P}[A_1|A_2]$$
 (exercise).

 $\rightsquigarrow$  Two examples...

(A) assume you bet b euros on "even" in one game of roulette...

 $\mathcal{E}$ : roulette,  $\Omega = \{0, 1, 2, ..., 36\}, \mathcal{A} = \mathcal{P}(\Omega), \mathbb{P}[A] = \frac{\#A}{\#\Omega}$ . Your (random) gain *X* is given by

$$X(\omega)=\left\{egin{array}{ccc} -b & ext{if} & \omega\in\{0,1,3,\ldots,35\}\ b & ext{if} & \omega\in\{2,4,\ldots,36\}, \end{array}
ight.$$

whose distribution is

distribution of X					
values	-b	b			
probabilities	19 37	$\frac{18}{37}$			

with expected winning  $-\frac{b}{37}$ .

(A) assume you bet b euros on "even" in one game of roulette...

$$\mathcal{E}$$
: roulette,  $\Omega = \{0, 1, 2, \dots, 36\}, \, \mathcal{A} = \mathcal{P}(\Omega), \, \mathbb{P}[\mathcal{A}] = rac{\#\mathcal{A}}{\#\Omega}$ 

Now, we are given the information that the event A="result is not 0" occurred? What is then your expected winning, namely  $\mathbb{E}[X|A]$ ? The r.v.  $X\mathbb{I}_A$  is determined by

$$(X\mathbb{I}_A)(\omega) = \left\{ egin{array}{ccc} -b & \mathrm{if} & \omega \in \{1,3,\ldots,35\} \\ 0 & \mathrm{if} & \omega = 0 \\ b & \mathrm{if} & \omega \in \{2,4,\ldots,36\}, \end{array} 
ight.$$

whose distribution is

distribution of $XI_A$				
values	-b	0	b	
probabilities	<u>18</u> 37	$\frac{1}{37}$	<u>18</u> 37	

Therefore,

your expected winning (conditional upon A) is

$$\mathbb{E}[X|A] = \frac{1}{\mathbb{P}[A]} \mathbb{E}[X\mathbb{I}_A] = \frac{1}{\left(\frac{36}{37}\right)} \left( (-b) \times \frac{18}{37} + 0 \times \frac{1}{37} + b \times \frac{18}{37} \right) = \mathbf{0}.$$

The fact that we know 0 won't be the result of the game, transforms this defavorable game ( $\mathbb{E}[X] = -b/37$ ) into a fair one.  $\mathbb{E}[X|A] = 0$ , i.e., the game is neither favorable nor defavorable.

(B) assume that three coins (10, 20, and 50 cents) are tossed and that you win the coins that show "head"...

 $\mathcal{E}$ : 3 coins,  $\Omega = \{HHH, HHT, \dots, TTT\}, \mathcal{A} = \mathcal{P}(\Omega), \mathbb{P}[A] = \frac{\#A}{\#\Omega}.$ 

Denote by X your gain or loss.

Denote by A the event "two heads and one tail",

that is,  $A = \{HHT, HTH, THH\}$ .

What is your conditional expected winning  $\mathbb{E}[X|A]$ ?

The r.v.  $XI_A$  is determined by

$$(X\mathbb{I}_{A})(\omega) = \begin{cases} 10+20 & \text{if } \omega = HHT\\ 10+50 & \text{if } \omega = HTH\\ 20+50 & \text{if } \omega = THH\\ 0 & \text{if } \omega \notin A. \end{cases}$$

The corresponding distribution is

distribution of $XI_A$					
values	30	60	70	0	
probabilities	1 8	18	1 8	58	

Therefore,

your expected winning (conditional upon A) is

$$\mathbb{E}[X|A] = \frac{\mathbb{E}[X\mathbb{I}_A]}{\mathbb{P}[A]} = \frac{1}{(\frac{3}{8})} \left( 30 \times \frac{1}{8} + \frac{60}{8} + \frac{70}{8} + 0 \times \frac{5}{8} \right) = 53.33...$$

This is to be compared with (exercise) the unconditional expected winning  $\mathbb{E}[X] = 40$  (interpretation).

Conditional expectation allows for exploiting some information (e.g., the information that some event occurred) in order to improve the (unconditional) best guess  $\mathbb{E}[X]$ .

This available information can take various forms:

- (the occurrence of) an event.
- (the value of ) a r.v.
- (the occurrence of some event in) a  $\sigma$ -algebra.

Let *X* be an integrable r.v. on  $(\Omega, \mathcal{A}, \mathbb{P})$ .

Let *Y* be a discrete r.v. on  $(\Omega, \mathcal{A}, \mathbb{P})$ , say with distribution

distribution of Y				
values	<i>Y</i> 1	<i>Y</i> 2		
probabilities	<i>p</i> <sub>1</sub>	p <sub>2</sub>		

Y partitions  $\Omega$  via

$$\Omega = \{\omega | Y(\omega) = y_1\} \cup \{\omega | Y(\omega) = y_2\} \cup \cdots$$

Then we define  $\mathbb{E}[X|Y]$  as the r.v.

 $\mathbb{E}[X|Y](\omega') = E[X|\{\omega|Y(\omega) = y_i\}] \quad \text{if} \quad \omega' \in \{\omega|Y(\omega) = y_i\}.$ 

The last conditional expectation is w.r.t. an event, and hence is well defined.

# **Conditional expectation**

Note: we can interpret  $\mathbb{E}[X|Y]$  as a function of  $\omega$  or as a function of *y*:

$$E[X|Y](\omega)$$
 or  $E[X|Y = y]$ .

It can be shown that:

#### Theorem:

(i)  $\mathbb{E}[X|Y]$  is  $\sigma(Y)$ -measurable [that is:  $\mathbb{E}[X|Y] = f(Y)$ ] (ii)  $\int_{A} \mathbb{E}[X|Y](\omega) dP(\omega) = \int_{A} X(\omega) dP(\omega)$ , for all  $A \in \sigma(Y)$ .

Actually, this characterizes the r.v.  $\mathbb{E}[X|Y]$ .

This will be used to extend the concept of conditional expectation to more general setups...

# **Conditional expectation**

Some remarks:

- ► If a r.v. Z is  $\sigma(Y)$ -measurable, then this means that Z = f(Y) for some (measurable) function f.
- If *E*[X<sup>2</sup>] is finite, then an alternative way to define *E*[X|Y] is this one:

 $E[X|Y] = f_0(Y)$ , with  $f_0$  is the function which minimizes the mean squared loss:

 $E[X-f_0(Y)]^2 \leq E[X-f(Y)]^2, \quad \forall f \text{ such that } E[f(Y)]^2 < \infty.$ 

We perform a non-linear regression!

Let *X* be an integrable r.v. on  $(\Omega, \mathcal{A}, P)$ . Then we define quite generally  $\mathbb{E}[X|Y]$  as <u>the</u> r.v. such that (i)  $\mathbb{E}[X|Y]$  is  $\sigma(Y)$ -measurable. (ii)  $\int_{\mathcal{A}} \mathbb{E}[X|Y](\omega) dP(\omega) = \int_{\mathcal{A}} X(\omega) dP(\omega)$ , for all  $\mathcal{A} \in \sigma(Y)$ .

In practice:

- If (X, Y) is discrete,  $\mathbb{E}[X|Y = y] = \sum_i x_i \mathbb{P}[X = x_i|Y = y]$ .
- If (X, Y) is absolutely continuous,

$$\mathbb{E}[X|Y = y] = \frac{1}{f^{Y}(y)} \int_{\mathbb{R}} x \, f^{(X,Y)}(x,y) \, dx = \int_{\mathbb{R}} x \, f^{(X|Y)}(x,y) \, dx,$$

where  $f^{(X|Y)}(x, y)$  denotes the pdf of X conditionally on Y = y.

Same properties:

•  $\mathbb{E}[X|Y]$  is  $\sigma(Y)$ -measurable

$$\blacktriangleright \mathbb{E}[X|Y] = \mathbb{E}[X] \text{ iff } X \perp Y.$$

•  $\mathbb{E}[X|Y] = X$  iff X is  $\sigma(Y)$ -measurable.

• 
$$\mathbb{E}\Big[\mathbb{E}[X|Y]\Big] = \mathbb{E}[X].$$

Remark:  $\mathbb{E}[X] = \mathbb{E}\left[\mathbb{E}[X|Y]\right] = \int_{\mathbb{R}} \mathbb{E}[X|Y = y] f^{Y}(y) dy$ . (// total probability formula).

We still define:  $\mathbb{P}[A|Y] = \mathbb{E}[\mathbb{I}_A|Y]].$ 

$$\sim \mathbb{P}[A] = \mathbb{E}[\mathbb{I}_A] = \mathbb{E}\left[\mathbb{E}[\mathbb{I}_A|Y]\right] = \int_{\mathbb{R}} \mathbb{E}[\mathbb{I}_A|Y = y] f^Y(y) \, dy = \int_{\mathbb{R}} \mathbb{P}[A|Y = y] f^Y(y) \, dy.$$

# **Conditional variances**

In time series analysis, many models are designed to explain the dynamics of the conditional variance  $Var[X_t|X_{t-1}, X_{t-2}, ...]$ (e.g., the so-called stochastic volatility models). We define

$$\operatorname{Var}[X|Y] = E[(X - E[X|Y])^2|Y].$$

Remarks:

Similarly (exercise) as for Var[X],

$$\operatorname{Var}[X|Y] = \mathbb{E}[X^2|Y] - (\mathbb{E}[X|Y])^2.$$

▶ But  $\operatorname{Var}[X] \neq \mathbb{E}[\operatorname{Var}[X|Y]]$ , rather (exercise) we have

$$\operatorname{Var}[X] = \mathbb{E}\left[\operatorname{Var}[X|Y]\right] + \operatorname{Var}\left[\mathbb{E}[X|Y]\right].$$

Example:

Let  $X \sim \text{Unif}(0, 1)$ . If X = x, you flip *m* times a coin with property  $\mathbb{P}[\text{Head}] = x$ . Let *N* be the number of "heads".  $\mathbb{E}[N] =$ ? Distribution of *N*?

$$\mathbb{E}[N|X = x] = \sum_{n=0}^{m} n \mathbb{P}[N = n|X = x] = \sum_{n=0}^{m} n \binom{m}{n} x^{n} (1 - x)^{m-n}$$
  
= mx expectation of a Binomial random variable.

 $\Rightarrow \mathbb{E}[N|X] = mX \ (\sigma(X) \text{-measurable}).$ 

Consequently,  $\mathbb{E}[N] = \mathbb{E}\left[\mathbb{E}[N|X]\right] = \mathbb{E}[mX] = m\mathbb{E}[X] = \frac{m}{2}$ .

# Distribution of N?

$$\mathbb{P}[N = n] = \int_{\mathbb{R}} \mathbb{P}[N = n | X = x] f^{X}(x) dx = \int_{0}^{1} \mathbb{P}[N = n | X = x] dx$$
  
=  $\int_{0}^{1} {m \choose n} x^{n} (1 - x)^{m - n} dx = {m \choose n} \int_{0}^{1} x^{n} (1 - x)^{m - n} dx$   
=  $\dots = \frac{1}{m + 1}$ ,

for all n = 0, 1, ..., m.

Hence, *N* is uniformly distributed on  $\{0, 1, ..., m\}$ . (which now makes clear why  $\mathbb{E}[N] = \frac{m}{2}$ ).

Example:

A hen lays *N* eggs, where  $N \sim \mathcal{P}(\lambda)$ . Each egg hatches out with probability *p*. Let *K* be number of born chicken.

 $\mathbb{E}[K|N] = ?$  and  $\mathbb{E}[K] = ?$ .

 $\mathbb{E}[K|N](\omega) = \mathbb{E}[K|N = n]$  if  $N(\omega) = n$ , where

$$\mathbb{E}[K|N=n] = \sum_{k} k \mathbb{P}[K=k|N=n] = \sum_{k=0}^{n} k \binom{n}{k} p^{k} q^{n-k}$$
$$= np \quad \text{expectation of a Binomial r.v.}$$

 $\Rightarrow \mathbb{E}[K|N] = Np \ (\sigma(N) \text{-measurable}).$ 

Consequently,  $\mathbb{E}[K] = \mathbb{E}[\mathbb{E}[K|N]] = \mathbb{E}[Np] = p \mathbb{E}[N] = p\lambda$ .

Example:

Let  $X \sim \text{Exp}(\lambda_1)$  and  $Y \sim \text{Exp}(\lambda_2)$  be two independent exponential random variables. What is the distribution of V = X + Y?

We use total probability formula and assume  $z \ge 0$ . (The case z < 0 is trivial.)

$$\begin{split} \mathbb{P}[V \leq z] &= \int_{\mathbb{R}} P[V \leq z | Y = y] f^{Y}(y) dy \\ &= \int_{\mathbb{R}} P[X + Y \leq z | Y = y] f^{Y}(y) dy \\ &= \int_{\mathbb{R}} P[X \leq z - y] f^{Y}(y) dy \\ &= \int_{\mathbb{R}} (1 - e^{-\lambda_{1}(z - y)}) \mathbb{I}[z - y \geq 0] \lambda_{2} e^{-\lambda_{2} y} \mathbb{I}[y \geq 0] dy \\ &= \int_{0}^{z} (1 - e^{-\lambda_{1}(z - y)}) \lambda_{2} e^{-\lambda_{2} y} dy = \cdots \end{split}$$

Conditional expectation allows for exploiting some information (e.g., the information that some event occurred) in order to improve the (unconditional) best guess  $\mathbb{E}[X]$ .

This available information can take various forms:

- (the occurrence of) an event.
- (the value of ) a r.v.
- (the occurrence of some event in) a  $\sigma$ -algebra.

# **Conditional expectation (w.r.t. a** $\sigma$ -algebra)

Let *X* be an integrable r.v. on  $(\Omega, \mathcal{A}, \mathbb{P})$ . Let *F* be a  $\sigma$ -algebra  $\subset \mathcal{A}$ .

Then we define  $\mathbb{E}[X|\mathcal{F}]$  as <u>the</u> r.v. such that (i)  $\mathbb{E}[X|\mathcal{F}]$  is  $\mathcal{F}$ -measurable. (ii)  $\int_{A} \mathbb{E}[X|\mathcal{F}](\omega) d\mathbb{P}(\omega) = \int_{A} X(\omega) d\mathbb{P}(\omega)$ , for all  $A \in \mathcal{F}$ .

Same kind of properties as for  $\mathbb{E}[X|Y]$ :

- $\mathbb{E}[X|\mathcal{F}]$  is  $\mathcal{F}$ -measurable
- $\blacktriangleright \mathbb{E}[X|\mathcal{F}] = \mathbb{E}[X] \text{ iff } X \perp \mathcal{F}.$
- $\mathbb{E}[X|\mathcal{F}] = X$  iff X is  $\mathcal{F}$ -measurable.

• 
$$\mathbb{E}\Big[\mathbb{E}[X|\mathcal{F}]\Big] = \mathbb{E}[X].$$

# **Conditional expectation (w.r.t. a** $\sigma$ -algebra)

Let  $\mathcal{F}_1 \subset \mathcal{F}_2$  be two  $\sigma$ -algebras in  $\mathcal{A}$ .

Extra properties:

• 
$$\mathbb{E}\left[\mathbb{E}[X|\mathcal{F}_2]|\mathcal{F}_1\right] = \mathbb{E}[X|\mathcal{F}_1].$$
  
•  $\mathbb{E}\left[\mathbb{E}[X|\mathcal{F}_1]|\mathcal{F}_2\right] = \mathbb{E}[X|\mathcal{F}_1].$ 

Conditional expectation w.r.t. to a  $\sigma$ -algebra also allows for defining conditional expectation w.r.t. a collection of r.v.'s. More specifically, we define

$$\mathbb{E}[X|Y_1,\ldots,Y_n] := \mathbb{E}[X|\sigma(Y_1,\ldots,Y_n)],$$

where  $\sigma(Y_1, \ldots, Y_n)$  is the smallest  $\sigma$ -algebra containing  $\{Y_i^{-1}(B) | B \in \mathcal{B}, i = 1, \ldots, n\}.$